

Utah State University
ECE 6010
Stochastic Processes
Homework # 7 Solutions

1. Suppose $\{X_t, t \geq 0\}$ is a homogeneous Poisson process with parameter λ . Define a random variable τ as the time of the first occurrence of an event. Find the p.d.f. and the mean of τ .

p.d.f. :

We have,

$$P(X_t = k) = \frac{e^{-\lambda t} (\lambda t)^k}{k!}$$

So, $P(\tau > t) = P(X_t = 0) = e^{-\lambda t}$
and $P(\tau \leq t) = 1 - e^{-\lambda t}$, i.e. $F_\tau(t) = 1 - e^{-\lambda t}$.

Therefore,

$$f_\tau(t) = \lambda e^{-\lambda t}$$

Mean :

$$E[\tau] = \int_0^\infty \tau \lambda e^{-\lambda \tau} d\tau = -\tau e^{-\lambda \tau} - \frac{1}{\lambda} e^{-\lambda \tau} \Big|_0^\infty = -0 - 0 + 0 + \frac{1}{\lambda} = \frac{1}{\lambda}$$

2. Suppose $\{X_t, t \in \mathbb{R}\}$ is a w.s.s. random process with autocorrelation function $R_X(\tau)$. Show that if R_X is continuous at $\tau = 0$ then it is continuous for all $\tau \in \mathbb{R}$.

$$\begin{aligned} [R_X(\tau + \delta) - R_X(\tau)]^2 &= (E[X_{\tau+\delta}X_0] - E[X_\tau X_0])^2 \\ &= (E[X_0(X_{\tau+\delta} - X_\tau)])^2 \\ &\leq E[X_0^2]E[(X_{\tau+\delta} - X_\tau)^2] \quad (\text{Cauchy-Schwartz}) \\ &= R_X(0)[R_X(0) - 2R_X(\delta) + R_X(0)] \\ &= 2R_X(0)[R_X(0) - R_X(\delta)] \\ &= \leq 2R_X(0)\varepsilon \end{aligned}$$

So,

$$\begin{aligned} \Rightarrow |R_X(\tau + \delta) - R_X(\tau)| &\leq \sqrt{2R_X(0)\varepsilon} \\ \Rightarrow |R_X(\tau + \delta) - R_X(\tau)| &< \varepsilon' \end{aligned}$$

So continuous at all τ .

3. Under the conditions of problem 2, show that for $a > 0$,

$$P(|X_{t+\tau} - X_t| \geq a) \leq \frac{2(R_X(0) - R_X(\tau))}{a^2}.$$

Let $g(x) = x^2$. g is non-negative, non-decreasing on $[0, \infty)$ and symmetric about 0. Then,

$$P(|X| \geq b) \leq \frac{E[g(x)]}{b^2}.$$

Now, our example corresponds to:

$$\begin{aligned}
 P(|X_{t+\tau} - X_t| \geq a) &\leq \frac{E[|X_{t+\tau} - X_t|^2]}{a^2} \\
 &= \frac{R_X(0) - 2R_X(\tau) + R_X(0)}{a^2} \\
 &= \frac{2(R_X(0) - R_X(\tau))}{a^2}
 \end{aligned}$$

4. Suppose A and B are random variables with $E[A^2] < \infty$ and $E[B^2] < \infty$. Define the random processes $\{X_t, t \in \mathbb{R}\}$ and $\{Y_t, t \in \mathbb{R}\}$ by

$$X_t = A + Bt \quad Y_t = B + At, \quad t \in \mathbb{R}.$$

Find the mean, autocorrelation, and cross correlations of these random processes in terms of the moments of A and B .

Mean :

$$\begin{aligned}
 \mu_X(t) &= E[A + Bt] = E[A] + tE[B] \\
 \mu_Y(t) &= E[B + At] = E[B] + tE[A]
 \end{aligned}$$

Autocorrelation :

$$\begin{aligned}
 R_X(t, s) &= E[X_t X_s] \\
 &= E[(A + Bt)(A + Bs)] \\
 &= E[A^2 + ABs + ABt + B^2ts] \\
 &= E[A^2] + E[AB](s + t) + E[B^2]ts
 \end{aligned}$$

Similarly,

$$R_Y(t, s) = E[B^2] + E[AB](s + t) + E[A^2]st$$

Cross correlation :

$$\begin{aligned}
 R_{XY}(t, s) &= E[X_t Y_s] \\
 &= E[(A + Bt)(B + As)] \\
 &= E[AB + A^2s + B^2t + ABts] \\
 &= E[AB](1 + st) + E[A^2]s + E[B^2]t
 \end{aligned}$$

5. Homogeneous Poisson

Simply take $p_k(t, s)$ and substitute it back into the differential equation and show that it works.

$$\begin{aligned}
 \frac{\partial}{\partial t} p_k(t, s) &= \frac{\partial}{\partial t} \frac{e^{-\lambda(t-s)} (\lambda(t-s))^k}{k!} \\
 &= \frac{-\lambda e^{-\lambda(t-s)} (\lambda(t-s))^k + e^{-\lambda(t-s)} k \lambda (\lambda(t-s))^{k-1}}{k!} \\
 &= \lambda \left[\frac{-e^{-\lambda(t-s)} (\lambda(t-s))^k}{k!} + \frac{e^{-\lambda(t-s)} (\lambda(t-s))^{k-1}}{(k-1)!} \right] \\
 &= \lambda [p_{k-1}(t, s) - p_k(t, s)]
 \end{aligned}$$

6. Inhomogeneous Poisson

Simply take $p_k(t, s)$ and substitute it back into the differential equation and show that it works.

$$\begin{aligned} \frac{\partial}{\partial t} p_k(t, s) &= \frac{-\lambda_t e^{-\int_s^t \lambda_x dx} + \lambda_t k (\int_s^t \lambda_x dx)^{k-1}}{k!} \\ &= \lambda_t \left[\frac{-e^{-\int_s^t \lambda_x dx}}{k!} + \frac{(\int_s^t \lambda_x dx)^{k-1}}{(k-1)!} \right] \\ &= \lambda_t [p_{k-1}(t, s) - p_k(t, s)] \end{aligned}$$

Mean: $\int_s^t \lambda_x dx$

Covariance: Assume that $t > s$:

$$\begin{aligned} E[X_t X_s] &= E[(X_t - X_s) X_s] + E[X_s^2] = E[X_t - X_s] E[X_s] + E[X_s^2] \\ &= \left[\int_0^t \lambda_x dx + 1 \right] \int_0^s \lambda_x dx \end{aligned}$$

Similarly when $t < s$. Then

$$R_X(t, s) = \left[\int_0^s \lambda_x dx + 1 \right] \int_0^t \lambda_x dx$$

7. Suppose $\{X_t, t \in \mathbb{R}\}$ is a random process with power spectral density

$$S_X(\omega) = \frac{1}{(1 + \omega^2)^2}.$$

Find the autocorrelation function of X_t .

$$\begin{aligned} R_X(\tau) &= \mathcal{F}^{-1}\{S_X(\omega)\} = \mathcal{F}^{-1}\left\{\frac{1}{(1 + \omega^2)}\right\} * \mathcal{F}^{-1}\left\{\frac{1}{(1 + \omega^2)}\right\} \\ &= \frac{1}{2} e^{-|\tau|} * \frac{1}{2} e^{-|\tau|} \\ &= \frac{1}{4} \int_{-\infty}^{\infty} e^{-|t|} e^{-|\tau-t|} dt \\ &= \frac{1}{4} \int_{-\infty}^0 e^t e^{t-\tau} dt + \frac{1}{4} \int_0^{\tau} e^{-t} e^{t-\tau} dt + \frac{1}{4} \int_{\tau}^{\infty} e^{-t} e^{\tau-t} dt \quad (\text{for } \tau \geq 0) \\ &= \frac{1}{4} \left[\int_{-\infty}^0 e^{2t-\tau} dt + \int_0^{\tau} e^{-\tau} dt + \int_{\tau}^{\infty} e^{t-2\tau} dt \right] \\ &= \frac{1}{4} \left[\frac{1}{2} e^{-\tau} + \tau e^{-\tau} + \frac{1}{2} e^{-\tau} \right] \\ &= \frac{1}{4} (\tau e^{-\tau} + e^{-\tau}) \quad (\text{for } \tau \geq 0) \end{aligned}$$

Similarly,

$$R_X(\tau) = \frac{1}{4} (-\tau e^{\tau} + e^{\tau}) \quad (\text{for } \tau < 0)$$

Therefore,

$$R_X(\tau) = \frac{1}{4} e^{-|\tau|} (|\tau| + 1)$$

8. Suppose that ω is a random variable with p.d.f. f_ω and θ is a random variable independent of ω uniformly distributed in $(-\pi, \pi)$. Define a random process by $X_t = a \cos(\omega t + \theta)$, $t \in \mathbb{R}$ where a is a constant. Find the power spectral density of $\{X_t\}$.

$$\begin{aligned}
 E[X_{t_1} X_{t_2}] &= E\{a^2 \cos(\omega t_1 + \theta) \cos(\omega t_2 + \theta)\} \\
 &= \frac{1}{2} a^2 E\{\cos(\omega t_1 - \omega t_2) - \cos(\omega t_1 + \omega t_2 + 2\theta)\} \\
 &= \frac{1}{2} a^2 \left[E\{\cos(\omega t_1 - \omega t_2)\} - \underbrace{E\{\cos(\omega t_1 + \omega t_2 + 2\theta)\}}_0 \right] \\
 &= \frac{1}{2} a^2 E\{\cos(\omega t_1 - \omega t_2)\} \\
 &= \frac{1}{2} a^2 \int_{-\infty}^{\infty} \cos(\tau\omega) f_\omega(\omega) d\omega = \frac{1}{2} a^2 \int_{-\infty}^{\infty} \frac{e^{j\omega\tau} + e^{-j\omega\tau}}{2} f_\omega(\omega) d\omega \\
 &= \frac{1}{4} a^2 2\pi [\mathcal{F}^{-1}\{f_\omega(\omega)\} + \mathcal{F}^{-1}\{f_\omega(-\omega)\}] \\
 &= \frac{\pi a^2}{2} [\mathcal{F}^{-1}\{f_\omega(\omega)\} + \mathcal{F}^{-1}\{f_\omega(-\omega)\}]
 \end{aligned}$$

Therefore,

$$S_X(\omega) = \frac{\pi a^2}{2} [f_\omega(\omega) + f_\omega(-\omega)]$$

9. Suppose that $\{X_t, t \in \mathbb{R}\}$ is a w.s.s., zero-mean, Gaussian random process with autocorrelation function $R_X(\tau), \tau \in \mathbb{R}$ and power spectral density $S_X(\omega), \omega \in \mathbb{R}$. Define the random process $\{Y_t, t \in \mathbb{R}\}$ by $Y_t = (X_t)^2, t \in \mathbb{R}$. find the mean, autocorrelation, and powerspectral density of $\{Y_t, t \in \mathbb{R}\}$.

Mean :

$$\mu_y(t) = E[X_t^2] = R_X(0) = \sigma^2$$

Autocorrelation :

$$\begin{aligned}
 R_Y(t, s) &= E[Y_t Y_s] = E[X_t^2 X_s^2] \\
 &= \frac{\partial^4}{\partial u^2 \partial v^2} \Phi_{X_t X_s}(u, v) \Big|_{u=v=0} \frac{1}{i^4} \\
 &\vdots \\
 &= R_X^2(0) + 2R_X^2(\tau) \quad (\tau = t - s)
 \end{aligned}$$

PSD :

$$S_Y(\omega) = \mathcal{F}\{R_Y(\tau)\} = 2\pi R_X^2(0)\delta(\omega) + 2(S_X(\omega) * S_X(\omega))$$

10. Suppose U and V are independent random variables with $E[U] = E[V] = 0$ and $\text{var}(U) = \text{var}(V) = 1$. Define random processes by

$$X_t = U \cos t + V \sin t \quad Y_t = U \sin t + V \cos t, \quad t \in \mathbb{R}.$$

Find the autocorrelation and cross-correlation functions of $\{X_t, t \in \mathbb{R}\}$ and $\{Y_t, t \in \mathbb{R}\}$. Are $\{X_t\}$ and $\{Y_t\}$ jointly wide sense stationary? Are they individually wide sense stationary?

$$\begin{aligned}
 R_X(t, s) &= E[X_t X_s] = E[(U \cos t + V \sin t)(U \cos s + V \sin s)] \\
 &= E[U^2 \cos t \cos s + UV(\cos t \sin s + \sin t \cos s) + V^2 \sin t \sin s] \\
 &= \cos t \cos s E[U^2] + E[U]E[V](\cos t \sin s + \sin t \cos s) + E[V^2] \sin t \sin s \\
 &= \cos t \cos s + \sin t \sin s \\
 &= \cos(t - s)
 \end{aligned}$$

Similarly,

$$R_Y(t, s) = \cos(t - s)$$

$$\begin{aligned}
 R_{XY}(t, s) &= E[X_t Y_s] = E[(U \cos t + V \sin t)(U \sin s + V \cos s)] \\
 &= E[U^2 \cos t \sin s + UV(\cos t \cos s + \sin t \sin s) + V^2 \sin t \cos s] \\
 &= E[U^2] \cos t \sin s + E[UV](\cos t \cos s + \sin t \sin s) + E[V^2] \sin t \cos s \\
 &= \cos t \sin s + \sin t \cos s \\
 &= \sin(t + s)
 \end{aligned}$$

$$\mu_X(t) = 0 \quad \mu_Y(t)$$

So, $\{X_t\}$ and $\{Y_t\}$ are individually WSS, but not jointly WSS.